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SAS Data Analysis Examples

Multinomial Logistic Regression

Version info: Code for this page was tested in SAS 9.3.

Multinomial logistic regression is for modeling nominal outcome variables, in which the log odds of the outcomes are modeled as a linear combination of the predictor variables.

Please Note: The purpose of this page is to show how to use various data analysis commands. It does not cover all aspects of the research process which researchers are expected to do. in particular, it does not cover data cleaning and checking, verification of assumptions, model diagnostics and potential follow-up analyses.

Examples of multinomial logistic regression

Example 1. People's occupational choices might be influenced by their parents' occupations and their own education level. We can study the relationship of one's occupation choice with education level and father's occupation. The occupational choices will be the outcome variable which consists of categories of occupations.

Example 2. A biologist may be interested in food choices that alligators make. Adult alligators might have difference preference than young ones. The outcome variable here will be the types of food, and the predictor variables might be the length of the alligators and other environmental variables.

Example 3. Entering high school students make program choices among general program, vocational program and academic program. Their choice might be modeled using their writing score and their social economic status.

Description of the data

For our data analysis example, we will expand the third example using the hsbdemo data set. You can download the data here.

```
proc contents data = "c:\hsbdemo";
run;
```

The CONTENTS Procedure

Data Set Name Member Type Engine Created Last Modified Protection Data Set Type	d:\data\hsbdemo DATA V9 Thursday, August 29, 2013 09:42:59 AM Thursday, August 29, 2013 09:42:59 AM	Observations Variables Indexes Observation Length Deleted Observation Compressed Sorted
Data Set Type	Writton by CAC	Sorted

Label Written by SAS

Data Representation WINDOWS_64
Encoding wlatin1 Western (Windows)

Engine/Host Dependent Information

Data Set Page Size	4096
Number of Data Set Pages	3
First Data Page	1
Max Obs per Page	101
Obs in First Data Page	42
Number of Data Set Repairs	0
Filename	d:\data\hsbdemo.sas7bdat
Release Created	9.0301M1
Host Created	X64_7PRO

Alphabetic List of Variables and Attributes

#	Variable	Туре	Len	Label
12	AWARDS	Num	3	
13	CID	Num	3	
2	FEMALE	Num	3	

11	HONORS	Num	3	honores eng
1	ID	Num	4	-
8	MATH	Num	3	math score
5	PROG	Num	3	type of program
6	READ	Num	3	reading score
4	SCHTYP	Num	3	type of school
9	SCIENCE	Num	, 3	science score
3	SES	Num	3	
10	SOCST	Num	3	social studies score
7	WRITE	Num	3	writing score

Sort Information

Sortedby PROG Validated YES Character Set ANSI

The data set contains variables on 200 students. The outcome variable is prog, program type. The predictor variables are social economic status, ses, a three-level categorical variable and writing score, write, a continuous variable. Let's start with getting some descriptive statistics of the variables of interest.

```
proc freq data = "c:\hsbdemo";
tables prog*ses / chisq norow nocol nofreq;
run;
```

The FREQ Procedure

Table of PROG by SES

PROG(type of program) SES

	1			
1	8.00	10.00	4.50	22.50
2	9.50	22.00	21.00	52.50
3	6.00	15.50	3.50	25.00
	47	95	-	200

Statistics for Table of PROG by SES

Prob
0.0023 0.0021 0.8068

Sample Size = 200

```
proc sort data = "c:\hsbdemo";
by prog;
run;
```

proc means data = "c:\hsbdemo";
var write;
by prog;
run;

type of program=1

The MEANS Procedure

Analysis Variable : WRITE writing score

N	Mean	Std Dev	Minimum	Maximum
45	51.3333333	9.3977754	31.0000000	67.0000000

type of program=2

Analysis Variable : WRITE writing score

N	Mean	Std Dev	Minimum	Maximum
105	56.2571429	7.9433433	33.0000000	67.0000000

type of program=3

Analysis Variable : WRITE writing score

Ŋ	Mean	Std Dev	Minimum	Maximum
50	46.7600000	9.3187544	31.0000000	67.0000000

Analysis methods you might consider

- · Multinomial logistic regression: the focus of this page.
- · Multinomial probit regression: similar to multinomial logistic regression but with independent normal error terms.
- · Multiple-group discriminant function analysis: A multivariate method for multinomial outcome variables
- Multiple logistic regression analyses, one for each pair of outcomes: One problem with this approach is that each analysis is potentially run on a
 different sample. The other problem is that without constraining the logistic models, we can end up with the probability of choosing all possible outcome
 categories greater than 1.
- Collapsing number of categories to two and then doing a logistic regression: This approach suffers from loss of information and changes the original
 research questions to very different ones.
- Ordinal logistic regression: If the outcome variable is truly ordered and if it also satisfies the assumption of proportional odds, then switching to ordinal logistic regression will make the model more parsimonious.
- Alternative-specific multinomial probit regression: allows different error structures therefore allows to relax the independence of irrelevant alternatives (IIA, see below "Things to Consider") assumption. This requires that the data structure be choice-specific.
- · Nested logit model: also relaxes the IIA assumption, also requires the data structure be choice-specific.

Multinomial logistic regression

Below we use proc logistic to estimate a multinomial logistic regression model. The outcome prog and the predictor ses are both categorical variables and should be indicated as such on the class statement. We can specify the baseline category for prog using (ref = "2") and the reference group for ses using (ref = "1"). The param=ref option on the class statement tells SAS to use dummy coding rather than effect coding for the variable ses.

```
proc logistic data = "c:\hsbdemo";
class prog (ref = "2") ses (ref = "1") / param = ref;
model prog = ses write / link = glogit;
run;
```

The LOGISTIC Procedure

Model Information

Data Set Response Variable Number of Response Levels Model Optimization Technique	d:\data\hsbdemo PROG 3 generalized logit Newton-Raphson	Written by SAS type of program
---	---	-----------------------------------

Number of Observations Read 200 Number of Observations Used 200

Response Profile

Ordered		Total
Value	PROG	Frequency

1	1	45
2	2	105
3	3	50

Logits modeled use PROG=2 as the reference category.

Class Level Information

Class	Value	Design Variables		
SES	1	0	0	
	2	1	0	
	3	0	1	

Model Convergence Status

Convergence criterion (GCONV=1E-8) satisfied.

Model Fit Statistics

		Intercept
	Intercept	and
Criterion	Only	Covariates
AIC	412.193	375.963
SC	418.790	402.350
-2 Log L	408.193	359.963

Testing Global Null Hypothesis: BETA=0

Test	Chi-Square	DF	Pr > ChiSq
Likelihood Ratio	48.2299	6	<.0001
Score	45.1588	6	<.0001
Wald	37.2946	6	<.0001

Type 3 Analysis of Effects

Effect	DF	Wald Chi-Square	Pr > ChiSq
SES	4	10.8162	0.0287
WRITE	2	26.4633	<.0001

Analysis of Maximum Likelihood Estimates

Parameter		PROG	DF	Estimate	Standard Error	Wald Chi-Square	Pr > ChiSq
Intercept		1	1	2.8522	1.1664	5.9790	0.0145
Intercept		3	1	5.2182	1.1635	20.1128	<.0001
SES	2	1	1	-0.5333	0.4437	1.4444	0.2294
SES	2	3	1	0.2914	0.4764	0.3742	0.5407
SES	3	1	1	-1.1628	0.5142	5.1137	0.0237
SES	3	3	1	-0.9827	0.5956	2.7224	0.0989
WRITE		1	1	-0.0579	0.0214	7.3200	0.0068
WRITE		3	1	-0.1136	0.0222	26.1392	<.0001

Odds Ratio Estimates

		Point	95% Wald
Effect	PROG	Estimate	Confidence Limits

SES	2	٧s	1	1	0.587	0.246	1.400
SEŞ	2	vs	1	3	1.338	0.526	3.404
SES	3	vs	1	1	0.313	0.114	0.856
SES	3	٧s	1	3	0.374	0.116	1.203
WRITE				1	0.944	0.905	0.984
WRITE				3	0.893	0.855	0.932

- In the output above, the likelihood ratio chi-square of 48.23 with a p-value < 0.0001 tells us that our model as a whole fits significantly better than an
 empty model (i.e., a model with no predictors)
 - · Several model fit measures such as the AIC are listed under Model Fit Statistics
- Two models are tested in this multinomial regression, one comparing membership to general versus academic program and one comparing membership to vocational versus academic program. They correspond to the two equations below:

$$ln\left(\frac{P(prog=general)}{P(prog=academic)}\right) = b_{10} + b_{11}(ses=2) + b_{12}(ses=3) + b_{13}write$$

$$ln\left(\frac{P(prog=vocation)}{P(prog=academic)}\right) = b_{20} + b_{21}(ses=2) + b_{22}(ses=3) + b_{23}write$$

where b 's are the regression coefficients.

- A one-unit increase in the variable write is associated with a .058 decrease in the relative log odds of being in general program vs. academic program.
- A one-unit increase in the variable write is associated with a .1136 decrease in the relative log odds of being in vocation program vs. academic
 program.
- The relative log odds of being in general program vs. in academic program will decrease by 1.163 if moving from the lowest level of ses (ses==1) to the highest level of ses (ses==3).
- · The overall effects of ses and write are listed under "Type 3 Analysis of Effects", and both are significant.
- The ratio of the probability of choosing one outcome category over the probability of choosing the baseline category is often referred to as relative risk (and it is also sometimes referred to as odds as we have just used to described the regression parameters above). Relative risk can be obtained by exponentiating the linear equations above, yielding regression coefficients that are relative risk ratios for a unit change in the predictor variable. In the case of two categories, relative risk ratios are equivalent to odds ratios, which are listed in the output as well.
 - The odds ratio for a one-unit increase in the variable write is .944 (exp(-.0579) from the regression coefficients above the odds ratios) for being in general program vs. academic program.
 - The odds ratio of switching from ses ≈ 1 to 3 is .313 for being in general program vs. academic program. In other words, the expected risk of staying in the general program is lower for subjects who are high in ses.

Using the test statement, we can also test specific hypotheses within or even across logits, such as if the effect of ses=3 in predicting general versus academic equals the effect of ses=3 in predicting vocational versus academic. Usage of the test statement requires the unique names SAS assigns each parameter in the model. The option outest = on the proc logistic statement produces an output dataset with the parameter names and values. We can get these names by printing them, and we transpose them to be more readable. The noobs option on the proc print statement suppresses observation numbers, since they are meaningless in the parameter dataset.

```
proc logistic data = "c:\hsbdemo" outest = mlogit param;
class prog (ref = "academic") ses (ref = "1") / param = ref;
model prog = ses write / link = glogit;
run:
proc transpose data = mlogit_param;
proc print noobs;
run;
NAME
               _LABEL
                                             PROG
Intercept 3
            Intercept: PROG=3
                                            2.546
Intercept_2 Intercept: PROG=2
                                          -1.689
SES1_3
             SES 1: PROG=3
                                          -0.180
SES1_2
              SES 1: PROG=2
                                          -1.163
SES2_3
SES2_2
             SES 2: PROG=3
SES 2: PROG=2
                                           0.645
                                          -0.630
              SES 3: PROG=3
SES3 3
                                           0.000
              SES 3: PROG=2
SES3 2
                                           0.000
WRITE 3
              writing score: PROG=3
                                          -0.056
```

Here we see the same parameters as in the output above, but with their unique SAS-given names. We are interested in testing whether SES3_general is equal to SES3_vocational, which we can now do with the test statement. The code

0.058

-179.982

writing score: PROG=2

Model Log Likelihood

WRITE 2

LNLIKE_

preceding the ":" on the test statement is a label identifying the test in the output, and it must conform to SAS variable-naming rules (i.e., 32 characters in length or less, letters, numerals, and underscore).

```
proc logistic data = "c:\hsbdemo" outest = mlogit_param;
class prog (ref = "2") ses (ref = "1") / param = ref;
model prog = ses write / link = glogit;
SES3_general_vs_SES3_vocational: test SES3_1 - SES3_3;
run:
```

SOME OUTPUT OMITTED

Linear Hypotheses Testing Results

Label	wald Chi-Square	DF	Pr > ChiSq
SES3 general vs SES3 vocational	0.0772	1	0.7811

The effect of ses=3 for predicting general versus academic is not different from the effect of ses=3 for predicting vocational versus academic. You can also use predicted probabilities to help you understand the model. You can calculate predicted probabilities using the Ismeans statement and the Illink option. For multinomial data, Ismeans requires glm rather than reference (dummy) coding, even though they are essentially the same, so be sure to respecify the coding on the class statement. However, glm coding only allows the last category to be the reference group (prog = vocational and ses = 3)and will ignore any other reference group specifications. Below we use Ismeans to calculate the predicted probability of choosing program type academic or general at each level of ses, holding write at its means.

```
proc logistic data = "c:\hsbdemo" outest = mlogit_param;
class prog ses / param = glm;
model prog = ses write / link = glogit;
lsmeans ses / e ilink cl;
run:
```

SOME OUTPUT OMITTED

Coefficients for SES Least Squares Means

Parameter	type of program	SES	Row1	Row2	Row3	Row4	Row!
Intercept	1		1	1	1		
Intercept	2					1	:
SES 1	1	1	1				
SES 1	2	1				1	
SES 2	1	2		1			
SES 2	2	2					
SES 3	1	3			1		
SES 3	2	3					
writing score	1		52.775	52.775	52.775		
writing score	2					52.775	52.77!

SOME OUTPUT OMITTED

SES Least Squares Means

type of program	SES	Mean	Standard Error of Mean	Lower Mean	Upper Mean
1 1 2 2 2	1 2 3 1 2	0.3582 0.2283 0.1785 0.4397 0.4777 0.7009	0.07264 0.04512 0.05405 0.07799 0.05526 0.06630	0.2158 0.1399 0.07256 0.2868 0.3694 0.5709	0.5006 0.3168 0.2844 0.5925 0.5861 0.8309

The predicted probabilities are in the "Mean" column. Thus, for **ses** = 3 and **write** = 52.775, we see that the probability of being the academic program (program type 2) is 0.1785; for the general program (program type 1), the probability is 0.7009. To obtain predicted probabilities for the program type vocational, we can reverse the ordering of the categories using the **descending** option on the **proc logistic** statement. This will make academic the reference group for **prog** and 3 the reference group for **ses**.

proc logistic data = "c:\hsbdemo" outest = mlogit_param descending;
class prog ses / param = glm;

model prog = ses write / link = glogit;
lsmeans ses / e ilink cl;
run;

SOME OUTPUT OMITTED

Coefficients for SES Least Squares Means

Parameter	type of program	SES	Row1	Row2	Row3	Row4	Row!
Intercept	3		1	1	1		
Intercept	2					1	· :
SES 1	3	1	1				
SES 1	2	1				1	
SES 2	3	2		1			
SES 2	2	2					:
SES 3	3	3			1		
SES 3	2	3					
writing score	3		52.775	52.775	52.775		
writing score	2					52.775	52.77!

SOME OUTPUT OMITTED

SES Least Squares Means

			Standard		
type of			Error of	Lower	Upper
program	SES	Mean	Mean	Mean	Mean
2	,	0 0001	0.000	0.00450	0.000
3	1	0.2021	0.05996	0.08459	0.3197
3	2	0.2939	0.05036	0.1952	0.3926
3	3	0.1206	0.04643	0.02960	0.2116
2	1	0.4397	0.07799	0.2868	0.5925
2	2	0.4777	0.05526	0.3694	0.5861
2	3	0.7009	0.06630	0.5709	0.8309

Here we see the probability of being in the vocational program when ses = 3 and write = 52.775 is 0.1206, which is what we would have expected since (1 - 0.7009 - 0.1785) = 0.1206, where 0.7009 and 0.1785 are the probabilities of being in the academic and general programs under the same conditions.

Things to consider

- The Independence of Irrelevant Alternatives (IIA) assumption: Roughly, the IIA assumption means that adding or deleting alternative outcome
 categories does not affect the odds among the remaining outcomes.
- Diagnostics and model fit: Unlike logistic regression where there are many statistics for performing model diagnostics, it is not as straightforward to do diagnostics with multinomial logistic regression models. Some model fit statistics are listed in the output.
- Pseudo-R-Squared: The R-squared offered in the output is basically the change in terms of log-likelihood from the intercept-only model to the current
 model. It does not convey the same information as the R-square for linear regression, even though it is still "the higher, the better".
- Sample size: Multinomial regression uses a maximum tikelihood estimation method. Therefore, it requires a large sample size. It also uses multiple
 equations. Therefore, it requires an even larger sample size than ordinal or binary logistic regression.
- Complete or quasi-complete separation: Complete separation implies that only one value of a predictor variable is associated with only one value of the
 response variable. You can tell from the output of the regression coefficients that something is wrong. You can then do a two-way tabulation of the
 outcome variable with the problematic variable to confirm this and then rerun the model without the problematic variable.
- Empty cells or small cells: You should check for empty or small cells by doing a crosstab between categorical predictors and the outcome variable. If a
 cell has very few cases (a small cell), the model may become unstable or it might not run at all.
- Sometimes observations are clustered into groups (e.g., people within families, students within classrooms). In such cases, you may want to see our
 page on non-independence within clusters.

See Also

· SAS Annotated Output: Multinomial Logistic Regression

References

- Hosmer, D. and Lemeshow, S. (2000) <u>Applied Logistic Regression (Second Edition)</u>. New York: John Wiley & Sons, Inc...
- Agresti, A. (1996) An Introduction to Categorical Data Analysis. New York: John Wiley & Sons, Inc.

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